

RVMs for classification

Lecture 25

Consider $K=2 : t \in \{0, 1\}$.

$$y(\vec{x}, \vec{w}) = G(\vec{w}^T \cdot \vec{\phi}(\vec{x}))$$

We will use a separate λ_i for each w_i .

Given $\vec{\lambda} = \underline{\lambda_1 \dots \lambda_M}$, we have:

$$\log p(\vec{w} | \vec{t}, \vec{\lambda}) = \underbrace{\sum_{n=1}^N}_{\text{posterior}} [t_n \log y_n + (1-t_n) \log(1-y_n)] - \underbrace{\log Z}_{\text{log Z}}$$

$$- \underbrace{\frac{1}{2} \vec{w}^T A \vec{w}}_{\text{log(prior)}} + \text{const}(\vec{w}) \quad (*)$$

$$A = \begin{pmatrix} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_M \end{pmatrix}$$

Maximize (*) w.r.t \vec{w} to get \vec{w}^* :

$$\frac{\partial}{\partial w_j} [t_n \log y_n + (1-t_n) \log(1-y_n)] =$$

$$= \left[\frac{t_n}{y_n} \psi_j(\vec{x}_n) - \frac{(1-t_n)}{1-y_n} \psi_j(\vec{x}_n) \right] = \frac{t_n(1-y_n) - (1-t_n)y_n}{y_n(1-y_n)} \psi_j(\vec{x}_n) \quad \approx$$

$$\underbrace{y_n(1-y_n)}_{\uparrow \text{from } G'(\vec{w}^T \vec{\phi})} \left(\psi_j(\vec{x}_n) (t_n - y_n) \right) = \varphi_{jn}^T (t_n - y_n).$$

$$\text{So, } \nabla_{\vec{w}} \log p(\vec{w} | \vec{t}, \vec{\lambda}) \Big|_{\vec{w}^*} = \varphi^T (\vec{t} - \vec{y}) - A \vec{w} \Big|_{\vec{w}^*} = 0, \text{ so that}$$

$$\vec{w}^* = A^{-1} \varphi^T (\vec{t} - \vec{y}).$$

Furthermore,

$$\nabla_{\vec{w}} \nabla_{\vec{w}} \log p(\vec{w} | \vec{t}, \vec{z}) = - \underbrace{\Phi^T B \Phi}_{N \times N} - A$$

$$B = \begin{pmatrix} y_1(1-y_1) & & & \\ & \ddots & & \\ & & \ddots & \\ & 0 & & y_N(1-y_N) \end{pmatrix}$$

$$p(\vec{w}^* | \vec{t}, \vec{z}) e^{-\frac{1}{2} (\vec{w} - \vec{w}^*)^T \Sigma^{-1} (\vec{w} - \vec{w}^*)}$$

Thus $p(\vec{w} | \vec{t}, \vec{z}) \approx \text{gaussian}$, such that

↑ Laplace appr'n
for posterior

$$\Sigma^{-1} = \Phi^T B \Phi + A, \quad \text{evaluated at } \vec{w}^*$$

evidence for \vec{z}

$$\begin{aligned} p(\vec{t} | \vec{z}) &= \int d\vec{w} p(\vec{t} | \vec{w}) p(\vec{w} | \vec{z}) \approx \\ &\approx p(\vec{t} | \vec{w}^*) p(\vec{w}^* | \vec{z}) (2\pi)^{M/2} |\Sigma|^{1/2}. \end{aligned}$$

Now we have

$$\begin{aligned} \frac{\partial}{\partial z_i} \log p(\vec{t} | \vec{z}) &= \frac{\partial}{\partial z_i} \log p(\vec{w}^* | \vec{z}) + \frac{1}{2} \frac{\partial}{\partial z_i} \log |\Sigma| = \\ &= \frac{\partial}{\partial z_i} \left[\underbrace{\frac{1}{2} \log |A|}_{\sum_{j=1}^M \log z_j} - \underbrace{\frac{1}{2} \vec{w}^{*T} A \vec{w}^*}_{\frac{\partial A}{\partial z_i} = \mathbb{I}_{(i,i)}} \right] - \frac{1}{2} \frac{\partial}{\partial z_i} \log |\Sigma^{-1}| = \\ &= \frac{1}{2z_i} - \frac{1}{2} (\vec{w}_i^*)^2 - \frac{1}{2} \text{Tr} \left(\underbrace{\Sigma}_{\frac{\partial \Sigma^{-1}}{\partial z_i}} \right) = \\ &= \frac{1}{2z_i} - \frac{1}{2} (\vec{w}_i^*)^2 - \frac{\sum_{ii}}{2} = 0. \quad \mathbb{I}_{(i,i)} \end{aligned}$$

Then $z_i(w^*)^2 = \underbrace{1 - d_i \sum_{j \neq i} \gamma_j}_{\text{"f_i"}}, \text{ or}$

$$d_i^{\text{new}} = \frac{\gamma_i}{(w_i^*)^2} \quad \begin{array}{l} \text{update} \\ \text{eq'n} \end{array}$$

$\underbrace{\quad}_{\text{evaluated at } \vec{d}^{\text{old}}, \vec{w}^*}$

o iterate to convergence between
estimating $\{\vec{w}^*\}$ & $\vec{d} \Rightarrow$ obtain
 $\{\Sigma\}$ sparse
solution,
usually much
sparser than SVMs

We can also analyze d_i dependence more
explicitly by using $\log p(\vec{t} | \vec{d}) \equiv L(\vec{d})$ &
splitting off the d_i -dependent terms:
 $L(\vec{d}) = L(\vec{d}_{-i}) + \lambda(d_i)$ as before

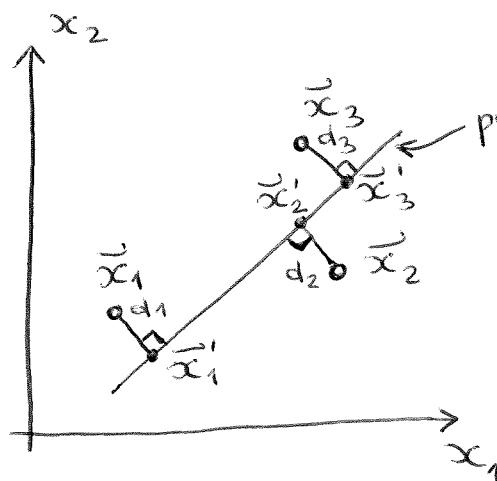
o Finally, RVMs generalize to $K > 2$
w/out difficulties:

$$y_k(\vec{x}) = \underbrace{\frac{e^{a_k}}{\sum_{j=1}^K e^{a_j}}}_{\text{softmax f'n}}, \quad a_k = \vec{w}_k^T \vec{x}$$

Principal component analysis (PCA)

goals: dimensionality reduction, feature extraction, data visualization

(D=2)



principal subspace

either maximize the variance of $\tilde{x}_1^1, \tilde{x}_2^1, \tilde{x}_3^1, \dots$ or minimize the distances d_1, d_2, d_3, \dots

Max variance

Consider $\{\tilde{x}_n\}_{n=1}^N$

formulation

dimensionality D

components
in \tilde{x}_n

goal: project \tilde{x}_n 's onto a subspace with $M < D$, while maximizing the variance of the projected data.

Start with $M=1$: define \tilde{u}_1 s.t. $\underbrace{\tilde{u}_1^\top \tilde{u}_1 = 1}_{D \text{ dim's}}$

Then $\tilde{x}_n \Rightarrow x_n^1 = \tilde{u}_1^\top \tilde{x}_n$, and

$$\langle \tilde{x}^1 \rangle = \underbrace{\frac{1}{N} \sum_{n=1}^N x_n^1}_{\text{mean}} = \tilde{u}_1^\top \cdot \underbrace{\frac{1}{N} \sum_n \tilde{x}_n}_{\langle \tilde{x} \rangle} = \tilde{u}_1^\top \cdot \langle \tilde{x} \rangle =$$

$$\text{Var}\{x_n^1\} = \frac{1}{N} \sum_n (\tilde{u}_1^\top \tilde{x}_n - \tilde{u}_1^\top \cdot \langle \tilde{x} \rangle)^2 =$$

$$= \frac{1}{N} \sum_n u_{1,k} [x_{n,k} x_{n,j} + \langle x \rangle_k \langle x \rangle_j - x_{n,k} \langle x \rangle_j - x_{n,j} \langle x \rangle_k] u_{1,j} \quad \text{①}$$

$$\textcircled{=} u_{1,k} \left\{ \underbrace{\frac{1}{N} \sum_n (x_{n,k} - \langle x \rangle_k) (x_{n,j} - \langle x \rangle_j)}_{\text{"S}_{kj}} \right\} u_{1,j} =$$

$= \tilde{u}_1^T S \tilde{u}_1$
 maximize this, subject to $\tilde{u}^T \tilde{u} = 1$:

$$J = \tilde{u}_1^T S \tilde{u}_1 + \lambda_1 (1 - \tilde{u}_1^T \tilde{u}_1)$$

\uparrow Lagrange multiplier

$$\frac{\partial J}{\partial \tilde{u}_1} = 2(S\tilde{u}_1 - \lambda_1 \tilde{u}_1) = 0 \Rightarrow S\tilde{u}_1 = \lambda_1 \tilde{u}_1$$

Thus \tilde{u}_1 is an eigenvector of S , with eigenvalue λ_1 .

Note that $\tilde{u}_1^T S \tilde{u}_1 = \lambda_1$, the variance will be at max if λ_1 is the largest eigenvalue of S ; corresponding \tilde{u}_1 is called the 1st principal component.

We can continue to project, s.t. for M-dim principal subspace the optimal projection is defined by M largest eigenvalues:

$$\lambda_1, \dots, \lambda_M$$

$\tilde{u}_1, \dots, \tilde{u}_M \Leftarrow$ corresponding eigenvectors

Proof. by induction:

M=1 shown above.

For $M > 1$, assume this holds for M & prove it for $M+1$. Thus

$\tilde{u}_1, \dots, \tilde{u}_M$ are principal eigenvectors
and we are looking for $\tilde{u}_{M+1} \perp \tilde{u}_1, \dots,$
 $\tilde{u}_{M+1} \perp \tilde{u}_M.$

$$\tilde{u}_{M+1}^T \tilde{u}_{M+1} = 1$$

otherwise \tilde{u}_{M+1} lies
in the M -subspace

Since Var in the \tilde{u}_{M+1} direction is
given by $\tilde{u}_{M+1}^T S \tilde{u}_{M+1}$, we need to
maximize

$$J = \tilde{u}_{M+1}^T S \tilde{u}_{M+1} + \lambda_{M+1} (1 - \tilde{u}_{M+1}^T \tilde{u}_{M+1}) + \\ + \sum_{i=1}^M \eta_i \tilde{u}_{M+1}^T \tilde{u}_i$$

$$\frac{\partial J}{\partial \tilde{u}_{M+1}} = 0 \Rightarrow 2S\tilde{u}_{M+1} - 2\lambda_{M+1}\tilde{u}_{M+1} + \sum_{i=1}^M \eta_i \tilde{u}_i = 0.$$

$$2\tilde{u}_{M+1}^T S \tilde{u}_{M+1} - 2\lambda_{M+1} \underbrace{\tilde{u}_{M+1}^T \tilde{u}_{M+1}}_1 + \sum_{i=1}^M \eta_i (\tilde{u}_{M+1}^T \tilde{u}_i) = 0, \text{ or}$$

$$\tilde{u}_{M+1}^T S \tilde{u}_{M+1} = \lambda_{M+1} \Leftrightarrow \begin{array}{l} \text{maximized} \\ \text{when } \lambda_{M+1} \text{ is} \end{array}$$

$$S\tilde{u}_{M+1} = \lambda_{M+1}\tilde{u}_{M+1}. \quad \begin{array}{l} \text{the largest} \\ \text{remaining} \\ \text{eigenvalue} \end{array}$$

Min-error formulation

again, introduce $\{\tilde{u}_i\}_{i=1}^D$ s.t.

$$\tilde{u}_j^\top \cdot \tilde{u}_i = \delta_{ij}$$

Then $\tilde{x}_n = \sum_{i=1}^D z_{ni} \tilde{u}_i = \sum_{i=1}^D (\tilde{x}_n^\top \cdot \tilde{u}_i) \tilde{u}_i$ indep. of n

approximate $\tilde{x}_n \Rightarrow \tilde{\tilde{x}}_n = \sum_{i=1}^M z_{ni} \tilde{u}_i + \sum_{i=M+1}^D b_i \tilde{u}_i$

Now, choose $\{\tilde{u}_i\}$, $\{z_{ni}\}$, $\{b_i\}$ to

minimize $J = \frac{1}{N} \sum_{n=1}^N \|\tilde{x}_n - \tilde{\tilde{x}}_n\|^2$

$$\frac{\partial J}{\partial z_{nj}} = \frac{2}{N} \left(\tilde{x}_n - \sum_{i=1}^M z_{ni} \tilde{u}_i - \sum_{i=M+1}^D b_i \tilde{u}_i \right)^\top \cdot \tilde{u}_j = \\ 1 \leq j \leq M$$

$$= \frac{2}{N} (\tilde{x}_n^\top \cdot \tilde{u}_j - z_{nj}) = 0 \Rightarrow z_{nj} = \tilde{x}_n^\top \cdot \tilde{u}_j$$

furthermore,

$$\frac{\partial J}{\partial b_i} = \frac{2}{N} \sum_{n=1}^N \left(\tilde{x}_n - \sum_{i=1}^M z_{ni} \tilde{u}_i - \sum_{i=M+1}^D b_i \tilde{u}_i \right)^\top \cdot \tilde{u}_j = \\ M+1 \leq j \leq D$$

$$= 2 \left(\frac{1}{N} \sum_{n=1}^N (\tilde{x}_n^\top \cdot \tilde{u}_j) - b_j \right) = 0, \text{ or}$$

$$b_j = \langle \tilde{x} \rangle^\top \cdot \tilde{u}_j$$

=

Now,

$$\tilde{x}_n - \tilde{\tilde{x}}_n = \sum_{i=M+1}^D [(\tilde{x}_n - \langle \tilde{x} \rangle)^T \cdot \tilde{u}_i] \tilde{u}_i$$

"residual" lies in the subspace defined by $\tilde{u}_{M+1}, \dots, \tilde{u}_D$.

Furthermore,

$$\begin{aligned} J &= \frac{1}{N} \sum_{n=1}^N \sum_{i=M+1}^D (\tilde{x}_n^T \cdot \tilde{u}_i - \langle \tilde{x} \rangle^T \cdot \tilde{u}_i)^2 = \\ &= \sum_{i=M+1}^D \tilde{u}_i^T S \tilde{u}_i. \end{aligned}$$

Consider $D=2, M=1$ case first:

minimize $J = \tilde{u}_2^T S \tilde{u}_2$ subject to $\tilde{u}_2^T \tilde{u}_2 = 1$.

$$\tilde{J} = \tilde{u}_2^T S \tilde{u}_2 + \lambda_2 (1 - \tilde{u}_2^T \tilde{u}_2).$$

$$\frac{\partial \tilde{J}}{\partial \tilde{u}_2} = 0 \Rightarrow S \tilde{u}_2 = \lambda_2 \tilde{u}_2, \Rightarrow J = \lambda_2$$

To minimize J , choose λ_2 to be the smaller eigenvalue \Rightarrow the principal subspace is defined by \tilde{u}_1 corresponding to λ_1 , the larger eigenvalue.

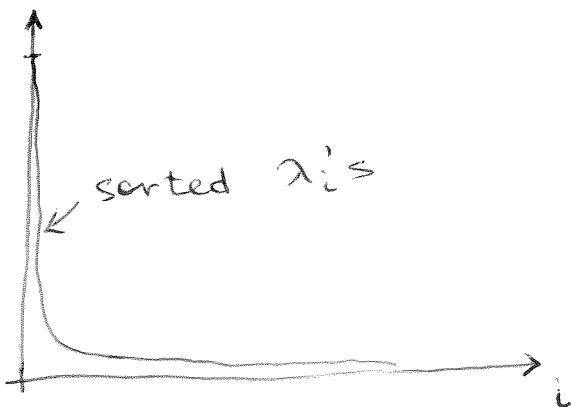
In general, for $M < D$ solve

$$\text{proof left as an exercise} \quad S\tilde{u}_i = \lambda_i \tilde{u}_i \quad i=1, \dots, D$$

Then $J = \sum_{i=M+1}^D \lambda_i$ is minimized by

choosing $\lambda_{M+1}, \dots, \lambda_D$ to be the smallest eigenvalues $\Rightarrow \lambda_1, \dots, \lambda_M$ are the largest eigenvalues & $\tilde{u}_1, \dots, \tilde{u}_M$ define the principal subspace.

Typically, we have for high- D data:



Clearly, $J \downarrow$ as $M \uparrow$

$$\text{Now, } \tilde{\bar{x}}_n = \sum_{i=1}^M (\tilde{x}_n^\top \cdot \tilde{u}_i) \tilde{u}_i + \sum_{i=M+1}^D (\tilde{x}_n^\top \cdot \tilde{u}_i) \tilde{u}_i \quad \textcircled{E}$$

$$\langle \tilde{x} \rangle = \sum_{i=1}^D (\langle \tilde{x} \rangle^\top \cdot \tilde{u}_i) \tilde{u}_i \quad \text{gives}$$

$$\textcircled{E} \quad \langle \tilde{x} \rangle + \underbrace{\sum_{i=1}^M (\tilde{x}_n^\top \cdot \tilde{u}_i - \langle \tilde{x} \rangle^\top \cdot \tilde{u}_i) \tilde{u}_i}_{\text{Components of an } M\text{-dim vector}}$$

This amounts to data compression

We can also use PCA to standardize the data.

Write $S\vec{u}_i = \lambda_i \vec{u}_i$ as $SU = UL$
each vector is a column

$$U = \begin{pmatrix} \vec{u}_1 & \dots & \vec{u}_D \end{pmatrix} \quad L = \begin{pmatrix} \lambda_1 & & 0 \\ 0 & \ddots & \lambda_D \end{pmatrix}$$

$$\underline{U^T = U^{-1}}$$

$$\text{Define } \vec{y}_n = L^{-1/2} U^T (\vec{x}_n - \langle \vec{x} \rangle).$$

Note that $\langle \vec{y} \rangle = \frac{1}{N} \sum_{n=1}^N \vec{y}_n = 0$, and covariance is then given by

$$\underbrace{\frac{1}{N} \sum_{n=1}^N \vec{y}_n \vec{y}_n^T}_{D \times D \text{ matrix}} = \frac{1}{N} \sum_{n=1}^N L^{-1/2} U^T (\vec{x}_n - \langle \vec{x} \rangle) (\vec{x}_n - \langle \vec{x} \rangle)^T * UL^{-1/2} \quad \textcircled{=}$$

$$\textcircled{=} \underbrace{L^{-1/2} U^T S U L^{-1/2}}_{U^{-1} U L = L} = \mathbb{I}.$$

This operation is known as whitening the data.

Finally, $M=2$ projections can be used for data visualization.

PCA for high-D data

Sometimes, $N < D \Rightarrow$ a set of N points occupy a subspace with at most $N-1$ dims, so using $M > N-1$ is since at least $D-N+1$ ~~dimensions~~ ~~are~~ ~~empty~~ ~~and~~ ~~all~~ ~~lambda's~~ are \emptyset . Doing "typical" PCA (i.e. finding all λ 's) scales as $\Theta(D^3)$ and is too costly.

To address this problem, define

$$X = \begin{pmatrix} \tilde{x}_1 - \langle \tilde{x} \rangle \\ \vdots \\ \tilde{x}_N - \langle \tilde{x} \rangle \end{pmatrix}$$

Then $\underbrace{S}_{\stackrel{D \times D}{N \times N}} = \frac{1}{N} X^T X$, and

$$\underbrace{\frac{1}{N} X^T X}_{S} \tilde{u}_i = \lambda_i \tilde{u}_i$$

eigenvector \Rightarrow $\underbrace{N^{-1} X X^T}_{N \times N} (\underbrace{X \tilde{u}_i}_{\vec{v}_i}) = \lambda_i (X \tilde{u}_i) \quad (*)$
 w/ut loss of generality assume $\vec{v}_i^T \vec{v}_i = 1$

$\lambda_i, i=1 \dots N$ are the same as the "original" λ 's, and the other λ 's are all \emptyset .

Solving $(*)$ is an $\Theta(N^3)$ operation.

Finally, eigenvectors:

$$(N^{-1} X^T X) \underbrace{(X^T \vec{v}_i)}_{\text{eigenvector of } S, \text{ length } D} = \lambda_i (X^T \vec{v}_i)$$

Normalization: $(X^T \vec{v}_i)^T (X^T \vec{v}_i) =$

$$= \vec{v}_i^T \underbrace{X X^T}_{N \lambda_i} \vec{v}_i = N \lambda_i \vec{v}_i^T \vec{v}_i$$

assuming that $\vec{v}_i^T \vec{v}_i = 1$, we obtain: $\frac{X^T \vec{v}_i}{\sqrt{N \lambda_i}}$ as normalized eigenvectors of S

Thus, we construct $N^{-1} X X^T$, compute its eigenvalues λ_i & eigenvectors \vec{v}_i ,
normalized length N

and then project back into D -space
using $\frac{X^T \vec{v}_i}{\sqrt{N \lambda_i}}$. \leftarrow normalized eigenvectors in original data space
length D