

Lecture 9 Bayesian information criterion (BIC)

Recall that under the Laplace approximation,

$$p(\vec{\theta} | \mathcal{D}, M) = \frac{p(\vec{\theta}, \mathcal{D} | M)}{\underbrace{p(\mathcal{D} | M)}_{\text{"Z from before"}}, \text{ where}$$

$$p(\mathcal{D} | M) = e^{\log p(\hat{\vec{\theta}}, \mathcal{D} | M)} (2\pi)^{D/2} |H|^{-1/2}$$

$D = \dim\{\vec{\theta}\}$

Using $\hat{\vec{\theta}} = \vec{\theta}_{\text{MAP}}$, we obtain:

$$\begin{aligned} \log p(\mathcal{D} | M) &= \underbrace{\log p(\mathcal{D} | \vec{\theta}_{\text{MAP}}, M)}_{\text{model type}} + \underbrace{\log p(\vec{\theta}_{\text{MAP}} | M)}_{\text{penalty terms}} - \frac{1}{2} \log |H| + \underbrace{\text{const}}_{\substack{\text{indep. of } M, \\ \text{indep. of } N}} \end{aligned}$$

log-likelihood

In the large- N limit, $\vec{\theta}_{\text{MAP}} \rightarrow \vec{\theta}^*$ MLE
and the effects of the prior
are negligible, such that

$$\log p(\mathcal{D} | M) = \log p(\mathcal{D} | \vec{\theta}^*, M) - \underbrace{\frac{1}{2} \log |H|}_{\text{Occam factor}}, \text{ where}$$

$$H_{ij} = \frac{\partial^2}{\partial \theta_i \partial \theta_j} \left[\underbrace{-\log p(\mathcal{D} | \vec{\theta}, M) - \log p(\vec{\theta} | M)}_{\substack{\text{discard} \\ \text{the prior}}} \right] \bigg|_{\vec{\theta}^*}$$

$$\textcircled{=} - \frac{\partial^2}{\partial \theta_i \partial \theta_j} \log p(\mathcal{D} | \vec{\theta}, M) \Big|_{\vec{\theta}^*}$$

$\underbrace{\sum_{k=1}^N \log p(\mathcal{D}_k | \vec{\theta}, M)}_{N \leftarrow \text{sample size}}$, additive over data points

We can write

$$H = \sum_{k=1}^N H_k = N \langle H \rangle, \text{ then}$$

$$\log |H| = \log |N \langle H \rangle| = \log(N^D \langle H \rangle) \textcircled{=}$$

$$\textcircled{=} D \log N + \underbrace{\log \langle H \rangle}_{\substack{\text{if } H \text{ is} \\ \text{full rank}}}$$

discard, indep. of N

Thus, $\underbrace{\log p(\mathcal{D} | M)}_{\text{BIC score}} \approx \log p(\mathcal{D} | \vec{\theta}^*, M) - \underbrace{\frac{D_M}{2} \log N}_{\text{model complexity penalty}}$

$D_M = \# \text{ parameters in model } M$

The issue of rank: suppose that ($D_M = 3$)

$$H_2 = \begin{pmatrix} H_{11} & H_{12} & 0 \\ H_{21} & H_{22} & 0 \\ 0 & 0 & 0 \end{pmatrix} = \begin{pmatrix} N \langle H_{11} \rangle & N \langle H_{12} \rangle & 0 \\ N \langle H_{21} \rangle & N \langle H_{22} \rangle & 0 \\ 0 & 0 & 0 \end{pmatrix}$$

due to likelihood only

and that $-\log p(\vec{\theta} | M) = \lambda |\vec{\theta}|^2$, then

$$H = H_{\mathcal{I}} + \lambda \mathbb{I} = \begin{pmatrix} N\langle H_{11} \rangle + \lambda & N\langle H_{12} \rangle & 0 \\ N\langle H_{21} \rangle & \underbrace{N\langle H_{22} \rangle + \lambda}_{\text{neglect}} & 0 \\ 0 & 0 & \underbrace{\lambda}_{\text{neglect}} \end{pmatrix}.$$

$$\begin{aligned} \text{Now, } \det H &= N^2 \langle H_{11} \rangle \langle H_{22} \rangle \lambda - N^2 \langle H_{12} \rangle \langle H_{21} \rangle \lambda = \\ &= N^2 \lambda [\langle H_{11} \rangle \langle H_{22} \rangle - \langle H_{12} \rangle \langle H_{21} \rangle]. \end{aligned}$$

Thus,

$$\log |H| = \underset{\substack{\uparrow \\ \text{not } 3!}}{2} \log N + \underbrace{\text{const}(N)}_{\substack{\text{discard} \\ \text{as before}}}$$

Information theory

Entropy of a discrete random variable X :

$$\begin{aligned} H(X) &= - \sum_{k=1}^K p(X=k) \log_2 p(X=k) = \\ &= -E[\log p(X)]. \end{aligned}$$

If $\log_2$ is used, units of entropy are bits.

Max entropy: ~~prob~~ $p(X=k) = \frac{1}{K}$, $\forall k$.

$$\text{Then } H(X) = \log_2 K.$$

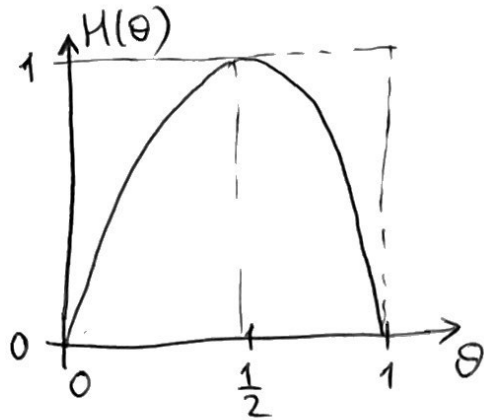
Min entropy: $p(X=k') = 1$, all others = 0.

$$\text{Then } H(X) = 0.$$

If $K=2$ (binary data), we have:

$$\begin{cases} p(X=1) = \theta, \\ p(X=0) = 1-\theta. \end{cases}$$

Then $H(X) = -\theta \log_2 \theta - (1-\theta) \log_2 (1-\theta) = H(\theta)$



$$\max\{H(\theta)\} = 1 \text{ bit}$$

Cross-entropy between discrete distributions P & Q :

$$H(p, q) = - \sum_{k=1}^K p_k \log_2 q_k$$

Joint entropy:

$$H(X, Y) = - \sum_{x, y} p(x, y) \log_2 p(x, y)$$

If X & Y are indep., $p(x, y) = p(x)p(y)$

and $H(X, Y) = H(X) + H(Y)$ upper bound on entropy

If X & Y are correlated, we expect

$$H(X, Y) < H(X) + H(Y)$$

Lower bound: $H(X, Y) \geq \max\{H(X), H(Y)\} \geq 0$. (*)

Conditional entropy:

$$\begin{aligned} H(Y|X) &= - \sum_x p(x) \sum_y p(y|x) \log p(y|x) = \\ &= - \sum_{x,y} p(x,y) \log \frac{p(x,y)}{p(x)} = - \sum_{x,y} p(x,y) \log p(x,y) + \\ &+ \sum_x p(x) \log p(x) = H(X,Y) - H(X). \end{aligned}$$

$$\text{Thus, } H(X,Y) = H(X) + H(Y|X).$$

$$\text{Likewise, } H(X,Y) = H(Y) + H(X|Y).$$

Since $H(X|Y) \geq 0$, $H(Y|X) \geq 0$, we recover Eq. (*).

moreover,

$$\underbrace{H(X,Y)}_{H(X) + H(Y|X)} \leq H(X) + H(Y) \Rightarrow H(Y|X) \leq H(Y)$$

If X completely determines Y : $H(Y|X) = 0$

If X & Y are indep.: $H(Y|X) = H(Y)$.

"Conditioning on data does not increase uncertainty, on average".

→ chain rule for entropy:

$$\begin{aligned} H(X_1, X_2, \dots, X_n) &= \sum_{i=1}^n H(X_i | X_1, \dots, X_{i-1}) = \\ &= H(X_1) + H(X_2 | X_1) + \dots + H(X_n | X_1, \dots, X_{n-1}) \end{aligned}$$

Continuous random variables:

Differential entropy:

$$h(X) = - \int dx p(x) \log p(x).$$

For example, if $X \sim \mathcal{U}(0, a)$,

$$h(X) = - \int_0^a dx \frac{1}{a} \log \frac{1}{a} = \log a,$$

can be < 0 if $a < 1$.

Similarly, if $X \sim \mathcal{N}(\mu, \sigma^2)$,

$$\begin{aligned} h(X) &= - \int dx p(x) \left[- (x - \mu)^2 / 2\sigma^2 - \frac{1}{2} \log(2\pi\sigma^2) \right] = \\ &= \frac{1}{2} + \frac{1}{2} \log(2\pi\sigma^2). \end{aligned}$$

KL divergence:

$$D_{KL}(p \parallel q) = \sum_{k=1}^K p_k \log \frac{p_k}{q_k} \quad \text{or}$$

$$D_{KL}(p \parallel q) = \int dx p(x) \log \frac{p(x)}{q(x)}.$$

Note that $D_{KL}(p \parallel q) = - \underbrace{\sum_k p_k \log q_k}_{H(p, q), \text{ cross entropy}} + \underbrace{\sum_k p_k \log p_k}_{-H(p), \text{ entropy of } p}$

Ex. $D_{KL}(p, q) = -H(p) - \int dx p(x) \left[-\frac{(x-\mu)^2}{2\sigma_2^2} - \frac{1}{2} \log(2\pi\sigma_2^2) \right]$ (1)

$$\begin{cases} p = \mathcal{N}(\mu, \sigma_1^2), \\ q = \mathcal{N}(\mu, \sigma_2^2). \end{cases}$$

$$\begin{aligned} &= -\frac{1}{2} - \frac{1}{2} \log(2\pi\sigma_1^2) + \frac{\sigma_1^2}{2\sigma_2^2} + \frac{1}{2} \log(2\pi\sigma_2^2) = \\ &= \log \frac{\sigma_2}{\sigma_1} + \frac{\sigma_1^2}{2\sigma_2^2} - \frac{1}{2}. \end{aligned}$$

If $\sigma_2 \gg \sigma_1$, $D_{KL} \approx \log \frac{\sigma_2}{\sigma_1} - \frac{1}{2} \geq 1$.

If $\sigma_2 \ll \sigma_1$, $D_{KL} \approx \frac{\sigma_1^2}{2\sigma_2^2} \gg 1$.

In fact, $D_{KL} \geq 0$, with $D_{KL} = 0$ iff $p = q$.

Recall Jensen's inequality:

for any concave function f ,

$$f\left(\sum_{i=1}^n \lambda_i \vec{x}_i\right) \leq \sum_{i=1}^n \lambda_i f(\vec{x}_i)$$

$$0 \leq \lambda_i \leq 1, \quad \sum_{i=1}^n \lambda_i = 1.$$

" f of the average is less than the average of f 's"

Now, consider $-D_{KL}(p \parallel q) = \sum_x p(x) \log \frac{q(x)}{p(x)} \leq$

$$\leq \log\left(\sum_x p(x) \frac{q(x)}{p(x)}\right) = \log 1 = 0.$$

↑
Jensen; log is concave

$$\Rightarrow \underline{D_{KL}(p \parallel q) \geq 0}.$$

